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Optimization of Culture Conditions for Hydrogen Production by an Anaerobic Bacteria Strain on Soluble Starch Using Response Surface Methodology

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Abstract

Bio-hydrogen is a clean source of energy with no harmful by-products produced during its combustion so hydrogen is potentially sustainable energy carrier for future. Therefore, bio-hydrogen produced by anaerobic bacteria in dark fermentation has attracted worldwide attention as renewable energy. However, capability of hydrogen production of these bacteria depends on major factors as substrates, iron-containing hydrogenase, reduction agent, pH and temperature. In this study, the Response Surface Methodology (RSM) with Central Composite Design (CCD) was employed to improve hydrogen production of a hydrogen-producing anaerobic bacteria strain that was isolated from animal waste in Phu Linh, Soc Son, Vietnam (PL strain). The hydrogen production process was investigated as a function of three critical factors: soluble starch concentration (8-12 g L-1), ferrous iron concentration (100-200 mg L-1) and L-cysteine concentration (300-500 mg L⁻¹). RSM analysis showed that all three factors had significant influences on the hydrogen production. Among them, ferrous iron concentration presented a greatest influence. The optimum hydrogen concentration of 1030 ml/L medium occurred with 10 g L⁻¹ of soluble starch, 150 mg L⁻¹ of ferrous iron and 400 mg L⁻¹ of L-cysteine after 48 hour- anaerobic fermentation. The hydrogen concentration that produced by the PL strain had increased to two times after using RSM. The obtained results indicated that RSM with CCD can be used as a technique to optimize culture conditions for enhancement of hydrogen production by the selected anaerobic bacteria strain. The production of hydrogen from low-cost organic substrates as soluble starch using anaerobic fermentation methods may be one of the most promising methods.

Keywords: Anaerobic bacteria strain; Bio-hydrogen; Dark fermentation; Response surface methodology

Introduction

Fossil fuels are not renewable and will be exhausted sooner or later. In addition, fossil fuel consumption caused global warming and environmental pollution. Thus, it is necessary to look for an alternative energy sources that are renewable and environmentally friendly. Hydrogen ($\rm H_2$) is an alternative energy source in the future because it has high specific energy content per unit mass (122 kJ/g) and produces no carbon-based emissions. Among various hydrogen production processes, biological methods are known to be less energy intensive than chemical or electrochemical ones since they are carried out at an ambient temperature and pressure [1,2].

However, one of the challenges of bio-hydrogen production processes is the feasibility of producing \mathbf{H}_2 at a commercial scale with a low cost to meet the need of sufficient and cost-effective energy supply. Consequently, the substrate used for fermentative \mathbf{H}_2 production must be abundant, easily available, and inexpensive. From those aspects, starch obtained from crop or food industry wastes could be a commercially viable Biohydrogen feedstock. In hydrogen production processes, dark fermentative hydrogen production from direct starch utilization was achieved using starch-fermenting bacteria, such as *Clostridium butyricum* [3].

Besides to substrate source, ferrous iron and L-cysteine are two important elements in dark fermentation hydrogen production. Ferrous iron is an essential element to form hydrogenase and ferredoxin, which is responsible for electrons transfer during hydrogen production process. L-cysteine is also an important nutrient for hydrogenase formation. In addition, L-cysteine, as a reducing agent, has the ability to reduce the Oxidation-Reduction Potential (ORP) value in the fermentation system, which enhances the growth of some hydrogen producing bacteria, thus considerably increasing the cumulative hydrogen production [4-

7]. Many studies have shown the effects of concentrations of soluble starch, ferrous iron and L-cysteine on hydrogen productivity based on the "one-variable-at-a-time" approach. However, these just carried out on individual factor but not mutual factors. So those results could not explain the mutual interactions among the independent variables and guarantee the determination of optimal conditions [4,5,8,9]. To the best of our knowledge, no study has ever investigated the interactions among these factors on the efficiency of the dark fermentation with soluble starch as substrate for hydrogen production. Therefore, the interactive effects of three key factors (soluble starch, ferrous iron and L-cysteine) were investigated in this study to maximize the hydrogen production by CCD of RSM [10-13].

The aim of this study was conducted to explore the effects of three independent factors (soluble starch, ferrous iron and L-cysteine) and interaction among them on $\rm H_2$ fermentation by an anaerobic bacteria strain (PL strain) using RSM, and find the optimum conditions for maximizing $\rm H_2$ production. The obtained results showed that the RSM with CCD was useful tool to finding the optimum condition for hydrogen producing by PL strain. The anaerobic bacteria (PL strain) had a great potential to produce $\rm H_2$ with soluble starch as substrate.

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Materials and Methods

Inoculum

The anaerobic hydrogen producing bacteria strain (PL strain) isolated from animal waste in Phu Linh village, Soc Son district, Hanoi, Vietnam was used as an inoculum.

Batch experiment set-up

The batch experiments were designed by RSM to study the effect of concentration of soluble starch, ferrous iron and L-cysteine on hydrogen production. The inoculum (5% v/v) was supplemented into 120 ml serum bottles with a working volume of 80 ml. The air in the head space of the serum bottles was removed by passing argon gas through the medium and the head space to provide anaerobic condition.

The medium used for fermentation consisted of (g L-¹): peptone 4; meat extracts 2; yeast extract 1; NaCl 2; $\rm K_2HPO_4$ 1.5; MgCl₂. 6H₂O 0.6; FeSO₄, 7H₂O 0.2; trace element 10 ml (MnSO₄, 7H₂O 0.01; ZnSO₄, 7H₂O 0.05; H₃BO₃ 0.01; CaCl₂. 2H₂O, 0.01; Na₂MoO₄ 0.01; CoCl₂. 6H₂O 0.2); vitamin solution 10 ml (riboflavin 0.025; citric acid 0.02; folic acid 0.01; para-aminobenzoic acid 0.01); soluble starch (6.95-13.05 g L-¹), ferrous iron (73.77-226.23 mg L-¹) and L-cysteine (247.54-552.46 mg L-¹) concentration were added into the medium according to various concentrations [13]. The initial pH of the medium was adjusted to 6.5 and the operation temperature was controlled at 30°C.

Full factorial central composite design for optimization of hydrogen production

Optimization studies with CCD were carried out by three variables including initial concentration of soluble starch, ferrous iron and L-cysteine (Tables 1 and 2). The hydrogen production was selected as the dependent output variable. For statistical calculations, the relation between the coded values and real values are described according to the following equation:

$$X_{i} = \frac{\left(X_{i} - X_{o}\right)}{\Delta X_{i}} \tag{1}$$

Where x_i is a coded value of the variable (Table 1) of the variable X_i , X_i is the real value of the i^{th} independent variable. X_o is the real value of X_i at the center point and ΔX_i is the step change of value. Initial soluble starch (X_1) , initial ferrous iron (X_2) , initial L-cysteine concentration (X_3) were chosen as three independent factors in the experiment design. A quadratic model was used to evaluate the optimization of key factor:

$$Y_{i} = \beta_{a} + \sum_{i} \beta_{i} x_{i} + \sum_{i} \beta_{i} x_{i}^{2} + \sum_{i} \beta_{i} X_{i} X_{i}$$
 (2)

Where Y_i is the predict response, β_o is constant, β_i is the linear coefficient, β_{ii} is the squared coefficient, and β_{ij} is the interaction coefficient.

The data analysis was calculated and analyzed using the "Design Expert" software (Version 7.1, Stat-Ease Inc., Minneapolis, USA). Statistical analysis of the model was performed to evaluate the analysis of variance (ANOVA) [14].

Analytical methods

Soluble starch concentration was measured by starch-iodide method. The substrate utilization was the ratio of starch consumption and initial amount of starch [15].

The evolved gas mixture in the headspace of the reactor was collected in a gas collector by displacement method at room temperature and atmospheric pressure. The quantity and composition of gas products (mainly, $\rm H_2$ and a little of $\rm CO_2$ and $\rm H_2S$) were determined in a gas

chromatograph GC-TCD (Thermo Trace GC-Thermo Electro-USA) equipped with a Thermal Conductivity Detector (TCD) and a column packed Molecular sieve 13×5 m. The operational temperatures of the oven and the detector were 50°C and 200°C, respectively. Heli was used as the carrier gas at a flow rate of 25 mL min⁻¹.

Results and Discussion

Levels of the independent variables and experimental ranges

Ranges of variables were chosen basically on investigating the effect of each variable on hydrogen production in batch experiments. Finally, levels of the independent variables were selected including: soluble starch concentration (8-12 g $L^{\text{-}1}$); ferrous iron concentration (100-200 mg $L^{\text{-}1}$); L-cysteine (300-500 mg $L^{\text{-}1}$). Each of the variables was coded at five levels: - α , -1, 0, +1 and + α (Table 1).

Based on CCD, a design of 20 experiments (Table 2) was formulated for three factorial (2³) design and six replicates at the central points, six star points were employed to the second-order polynomial model. The optimum values of the selected variables were obtained by solving regression equation and also by analyzing the response surface spots (Equation 3).

Optimization of the key factors for hydrogen production by PL strain

Soluble starch, ferrous iron and L-cysteine selected as key parameters were investigated the interactive effects to maximize the hydrogen production by CCD of RSM. By applying multiple regression analysis on the experimental data (Table 2), the following second order polynomial equation was established to explain the hydrogen production:

$$Y_{H2} = -7243.21 + 617.38 * X_{1} + 19.61 * X_{2} + 16.99 * X_{3} - 0.11 * X_{1}$$

$$*X_{2} - 0.23 * X_{1} * X_{3} - 0.002 * X_{2} * X_{3} - 23.98 * X_{1}^{2} - 0.06 * X_{2}^{2} - 0.018 * X_{3}^{2}$$
(3)

Where Y is the predicted hydrogen yield; X_1 , X_2 and X_3 are the coded values of soluble starch, ferrous iron and L-cysteine, respectively.

The analysis of variance (ANOVA) was conducted to test the significance of the fit of the second-order polynomial equation for the experimental data as shown in Table 3. The P-values are used as a tool to check significance of each variable, which also indicate the interaction strength between each independent variable. The smaller P-values, the bigger the significance of the corresponding variable [13].

The Model F-value of 507.67 implies the model is significant. There is only less than 99.99% (*p*-value<0.0001) chance that a "Model F-value" could occur due to noise.

The "Lack of Fit F-value" of 0.93 implies the "Lack of Fit" is not significant relative to the pure error. There is a 52.85% chance that a "Lack of Fit F-value" this large could occur due to noise. Non-significant lack of fit is good.

Here the R-squared value was 99.78%, which could explain 99.78% variability of the response. It indicates a good agreement between experiments and predicted values. It implies that the mathematical model is very reliable for hydrogen production in the present study.

The "Pred R-squared" of 98.99% is in reasonable agreement with the "Adj R-Squared" of 99.59%. "Adeq Precision" measures the signal to noise ratio. A ratio greater than 4 is desirable. The ratio of 68.109 indicates an adequate signal. This model can be used to navigate the design space [14] (Table 3).

Independent variables	Range and level coded value (x _i)						
	-α	-1	0	+1	+α		
X ₁ : initial soluble starch (g L ⁻¹)	6.95	8	10	12	13.05		
X ₂ : initial ferrous iron (mg L ⁻¹)	73.77	100	150	200	226.23		
X ₃ : initial L-cysteine (mg L ⁻¹)	247.54	300	400	500	552.46		

Table 1: Experimental range and levels of the independent variables.

Run			Independ	ent values	Hydrogen production (ml L-1)		Categorical factor levels		
		Coded values		Real values			Observed	Predicted	
No.	X ₁	X ₂	X ₃	X ₁	X ₂	X ₃			
1	+1	+1	+1	12	200	500	650	644.43	Fractional 2 ³ fractional factorial points
2	+1	+1	-1	12	200	300	700	702	
3	+1	-1	-1	12	100	300	600	599.63	
4	+1	-1	+1	12	100	500	575	584.4	
5	-1	+1	-1	8	200	300	520	517.56	
6	-1	+1	+1	8	200	500	640	647.43	
7	-1	-1	-1	8	100	300	360	372.53	
8	-1	-1	+1	8	100	500	540	544.81	
9	+α	0	0	13.05	150	400	850	849.46	Star point (6 points)
10	- α	0	0	6.95	150	400	690	678.56	
11	0	+α	0	10	226.23	400	540	535.35	
12	0	-α	0	10	73.76	400	600	585.85	
13	0	0	+α	10	150	552.46	630	622.67	
14	0	0	-α	10	150	247.54	540	544.81	
15	0	0	0	10	150	400	960	987.02	Central points (6 points)
16	0	0	0	10	150	400	992	987.02	
17	0	0	0	10	150	400	990	987.02	
18	0	0	0	10	150	400	993	987.02	
19	0	0	0	10	150	400	994	987.02	
20	0	0	0	10	150	400	993	987.02	

 Table 2: Central composite experimental design with three independent variables.

Effect interaction among of three factors on hydrogen production

Response surface plots are shown in Figure 1, which depict the interactions between two variables by keeping the other variables at their zero levels for hydrogen production.

Effect of interaction between soluble starch and ferrous iron

Interaction effect of soluble starch and ferrous iron on hydrogen production was displayed by the response surface plot in Figure 1a and Tables 2 and 3 with concentration of soluble starch and ferrous iron range from 8 to 12 g L⁻¹, 100 to 200 mg L⁻¹, respectively. *P*-values of each factor (soluble starch and ferrous iron) less than 0.05 indicated these factors all impact on hydrogen production. The hydrogen production was also influence by the interaction between two factors because *p*-valuable (p<0.0403) was less than 0.05 (Table 2).

Figure 1a and Table 2 show that hydrogen production decrease when soluble starch and ferrous iron concentration at level -1 (8 g $L^{\text{-1}}$ with soluble starch and 100 mg $L^{\text{-1}}$ with ferrous iron), gas mixture volume obtained about 360 ml $L^{\text{-1}}$. With higher soluble starch concentration at level +1 (12 g $L^{\text{-1}}$), gas mixture yield increase from 575 ml $L^{\text{-1}}$ to 700 ml $L^{\text{-1}}$ when ferrous iron concentration changes at two level -1 and +1 (100 mg $L^{\text{-1}}$ and 200 mg $L^{\text{-1}}$).

At central points, level 0 with 10 g $L^{\text{-}1}$ soluble starch, total gas yield was about 600 and 540 ml $L^{\text{-}1}$ with ferrous iron at level $-\alpha$ (73.76 mg $L^{\text{-}1}$) and level $+\alpha$ (226.23 mg $L^{\text{-}1}$), respectively. Although ferrous iron concentration at level $+\alpha$ was so higher than that at level $-\alpha$, total gas yield has an approximate value. It could be superfluous ferrous iron showed slightly inhibitive influence on hydrogen production in this study.

Maximum total gas production obtained approximately 995 ml L⁻¹ with 10 g soluble starch L⁻¹ and 150 mg ferrous iron L⁻¹. The concentration of ferrous iron and soluble starch used in this study were lower than that of other studies [2,3]. Consequently, external iron addition could improve bio-hydrogen production since hydrogenase, key enzyme on hydrogen production, contains iron at the active site. Increasing of iron concentration improve hydrogenase activity thus enhanced bio-hydrogen production. Nonetheless, too much iron concentration may be toxic for hydrogen-producing microorganisms [9,15]. Optimum concentration of soluble starch and ferrous iron for hydrogen production in this study was the same results reported by Yang et al. [16].

Effect of interaction among soluble starch and L-cysteine

Interaction between two variables (soluble starch and L-cysteine) on hydrogen production was displayed by the response surface plot

Source	Sum of Squares	df	Mean Square	F Value	p-value Prob>F	
Model	743801.02	9	82644.55	507.67	<0.0001	
X ₁ - Soluble starch	39734.13	1	39734.13	244.078	<0.0001	
X ₂ -Ferrous iron	33242.54	1	33242.54	204.2	<0.0001	
X ₃ -L-cystein	10372.44	1	10372.44	63.72	<0.0001	
X ₁ X ₂	903.13	1	903.13	5.55	0.0403	
X ₁ X ₃	17578.13	1	17578.13	107.98	<0.0001	
X_2X_3	903.13	1	903.13	5.55	0.0403	
X ₁ ^2	99463.50	1	99463.50	610.98	<0.0001	
X ₂ ^2	208665.97	1	208665.96	1281.79	<0.0001	
X ₃ ^2	332938.06	1	332938.06	2045.16	<0.0001	
Residual	1627.93	10	162.79			
Lack of Fit	786.59	5	157.32	0.93	0.5285	
Pure Error	841.33	5	168.27		<0.0001	
Cor Total	745428.95	19			<0.0001	
R-Squared: 0.9978	Adj R-Squared: 0.9959		Pred R-Squared: 0.9899	Adeq Precission: 68.109		

 Table 3: ANOVA for response surface quadratic model for hydrogen production.

in Figure 1b and Table 3. *P*-value of two factors (p_{XIX3} <0.0001) in Table 3 showed that interaction of those has significant to hydrogen production. This interaction between soluble starch and L-cysteine was stronger effect than ferrous iron and L-cysteine. Taherdanak et al. used CCD to find optimum factor as starch, ferrous iron and nickel via mesophilic dark fermentation. Results showed that starch and ferrous iron concentration were found to be the most important factor affecting the hydrogen production [8]. The different results might be explained that hydrogen was produced by mixed culture not pure culture. Table 2 and Figure 1b show that soluble starch concentration at level +1 (12 g L⁻¹), total gas yield obtained from 575 to 700 ml L⁻¹ with L-cysteine concentration at level -1 and +1. Total gas yield increased when starch concentration (8 g L⁻¹) and L-cysteine (100 mg L⁻¹) at level -1.

At central points, with 10 g soluble starch L⁻¹ and 400 mg L-cysteine L⁻¹, maximum total gas achieved above 990 ml L⁻¹. According to Houqing et al., maximum hydrogen production obtained by *Enterobacterium bacterium* M580 when 300 mg L-cysteine L⁻¹ was added into the culture with substrate as glucose [4]. Yuan et al. investigated the effect of L-cysteine on dark fermentative H₂ production and demonstrated that the dark fermentative hydrogen production using anaerobic mixed cultures was significantly enhanced by adding a small amount of L-cysteine (0.1-1.0 mM) as nutrient solution. The authors showed that optimum L-cysteine was 0.6 mM (72.69 mg L-cysteine L⁻¹) for maximum hydrogen production on dark fermentation on substrate as sucrose [17].

Effect of interaction among ferrous iron and L-cysteine

Interaction between two variables (ferrous iron and L-cysteine) on hydrogen production was displayed by the response surface plot in Figure 1c and Table 3. P-value of each variable less than 0.0001 showed the important role of these factors on hydrogen production. In addition, P-value of X_2X_3 less than 0.05 demonstrated that interaction between two variables had effect to hydrogen production (Table 3).

The response surface plot in Figure 1c showed the interactive effect between ferrous iron and L-cysteine when soluble starch concentration at level 0. When ferrous iron concentration at level-1 (100 mg $L^{\text{-}1}$), total gas production all decreased under 600 ml $L^{\text{-}1}$ at level -1 (300 mg $L^{\text{-}1}$) and +1 (500 mg $L^{\text{-}1}$) L-cysteine concentration. At level +1, - α of ferrous concentration, total gas obtained 600 - 650 ml $L^{\text{-}1}$ at level +1, -1 and 0. Al level + α ferrous iron and level 0 L-cysteine, hydrogen yield obtained

540 ml L^{-1} . Maximum total gas production achieved above 990 ml L^{-1} at 150 mg L^{-1} ferrous iron and 400 mg L-cysteine L^{-1} .

Yang et al. indicated that the highest gas yield at initial iron concentration lower than 100 mg Fe²⁺ L⁻¹. Initial iron concentration 55.3 mg Fe²⁺ L⁻¹ was reported producing maximum hydrogen yield from soluble starch by mixed anaerobes under mesophilic temperature [16]. The highest bio-hydrogen production obtained from sucrose by mixed microbes under mesophilic condition when initial iron concentration 73.47 mg Fe²⁺ L⁻¹ [15]. Whereas, some studies reported optimal initial iron concentration higher than 300 mg Fe²⁺ L⁻¹, initial ferrous concentration 300 mg Fe²⁺ L⁻¹ was optimal for hydrogen production from sucrose [11]. Furthermore, the highest hydrogen yield from sucrose was found at initial iron concentration 587.76 mg Fe²⁺ L⁻¹ [2]. Bao et al. reported that adding trace elements (Fe²⁺) and L-cysteine the hydrogen yields were higher than for control [5].

Fermentation under the optimal conditions

The maximum total gas production was determined about 1030 ml L^{-1} by calculating via CCD. The optimum condition for total gas producing by PL strain was 10 g soluble starch L^{-1} , 150 mg Fe²⁺ L^{-1} and 400 mg L^{-1} cysteine L^{-1} .

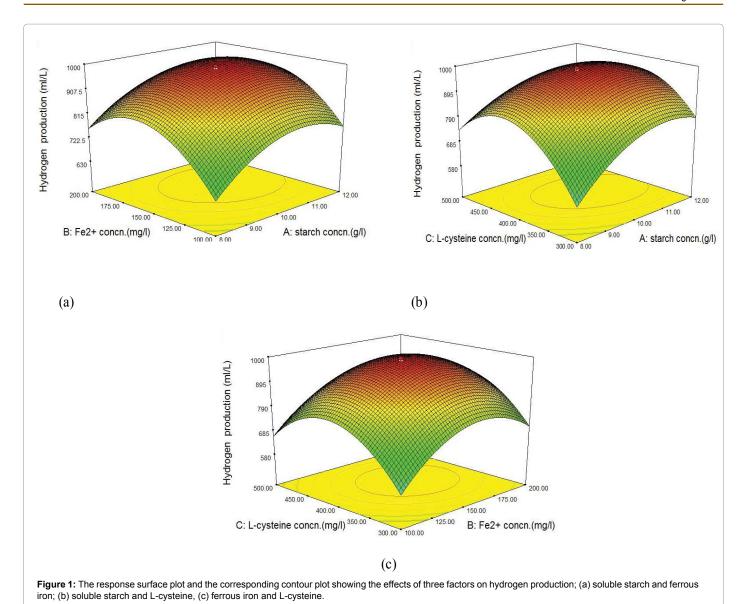
This strain was fermented in basal medium with 600 ml by adding 10 g soluble starch L⁻¹, 150 mg Fe²⁺ L⁻¹ and 400 mg L-cysteine L⁻¹ in fact. In optimum condition, PL strain consumed 98% soluble starch to produce 1030 ml total gas/L medium. The results of gas analysis by GC-TCD showed that a maximum hydrogen production of 91.85 ml H₂/g starch was achieved under optimum condition because H₂ gas occupied a volume of 91%. Similar result was obtained by Zhang et al. The author reported that 92 ml H₂/g soluble starch was produced by the *Thermoanaerobacterium* under thermophilic condition [18].

The obtained hydrogen volume increased to two times in comparison with previous studies that investigated the suitable conditions on hydrogen production with 10 g starch L^{-1} , 100 mg Fe^{2+} L^{-1} and 500 mg L-cysteine L^{-1} (data not shown) [19].

RSM was used to optimize critical factors for hydrogen fermentation by bacteria in many studies. Using RSM, Guo et al. indicated that the most important factor for hydrogen producing by *Ethanoligenens harbinense* 49 were glucose [13], Fe²+ and Mg²+. Maximum H₂ yield obtained in optimum condition was 2.20 mol/ mol glucose [10]. Yang et al. determined the important factors for hydrogen production by an anaerobic was temperature, pH and glucose. The optimum value of these factors was detected by RSM. At this condition, maximum H₂ yield obtained 1.75 mol/ mol glucose [9].

However, to best of our knowledge, there was no study using RSM to optimize these three factors (starch, Fe²⁺ and L-cysteine) together. In this study, soluble starch, Fe²⁺ and L-cysteine played important roles for hydrogen production by anaerobic PL strain. In fact that starch was a substrate for bacteria growth. Fe²⁺ is an essential element to form key enzyme on hydrogen production, increasing of iron improve hydrogenase activity thus enhanced biohydrogen production [20]. L-cysteine was reported to reduce OPR value of fermentation system. L-cysteine was also claimed to be a mediator between bacteria and substrate [17]. These reasons explained that three selected factors were necessary and all had significant affect to hydrogen production by anaerobic PL strain in our study.

The maximum $\rm H_2$ yield-91.85 ml/g starch (3.71 mol $\rm H_2$ / g starch) obtained in the optimum condition higher than results in studies of Guo et al. (2.20 mol $\rm H_2$ / mol glucose) and Mu et al. (1.75 mol $\rm H_2$ / mol glucose) [9,10]. The reason was probably that adding amount of



L-cysteine increase the utility rate of the substrate resulting enhancing maximum hydrogen yield [21].

Conclusion

This study focused on the optimization of the key factors for enhancement the bio-hydrogen production using the statistical methodology. Experimental results showed that three independent variables (soluble starch, ferrous iron and L-cysteine) and interaction between them had significant influences on the hydrogen production potential. The optimum conditions for fermentation hydrogen were basal medium with 10 g soluble starch L⁻¹, 150 mg Fe²⁺ L⁻¹, 400 mg L-cysteine L⁻¹. In this optimum condition, maximum H₂ gas obtained 91.85 ml H₂ /g starch (1030 ml total gas/L medium). The obtained results showed that the RSM was useful to optimize the hydrogen production process and to improve the hydrogen production potential by PL strain isolated in Vietnam.

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